



# Do men consult less than women? An analysis of routinely collected UK general practice data

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**Title Page**

**Do men consult less than women? An analysis of routinely collected UK general practice data**

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## Abstract

**Objective:** To examine whether gender differences in primary care consultation rates a) vary by age and deprivation status, and b) diminish when consultation for reproductive reasons or common underlying morbidities are accounted for.

**Design:** Cross sectional study of a cohort of patients registered with general practice.

**Setting:** UK primary care

**Subjects:** Patients (1,869,149 men, 1,916,898 women) registered with 446 eligible practices in 2010

**Primary outcome measures:** Primary care consultation rate

**Results:** This study analyses routinely collected primary care consultation data. The crude consultation rate was 32% lower in men than women. The magnitude of gender difference varied across the life-course, and there was no “excess” female consulting in early and later life. The greatest gender gap in primary care consultations was seen amongst those aged between 16 and 60 years. Gender differences in consulting were higher in people from more deprived areas than amongst those from more affluent areas. Accounting for reproductive related consultations diminished but did not eradicate the gender gap. However, consultation rates in men and women who had comparable underlying morbidities were similar; men with depression were only 8% less likely to consult than women with depression (relative risk 0.916, 95% confidence interval (CI) 0.913-0.918), and men with cardiovascular disease were just 5% less likely to consult (RR=0.950, 95%CI 0.948-0.952) than women with cardiovascular disease. These small gender differences diminished further, particularly for depression (RR=0.950, 95% CI 0.947-0.953), after also taking account of reproductive consultations.

**Conclusions:** Overall gender differences in consulting are most marked between the ages of 16 and 60 years; these differences are only partially accounted for by consultations for reproductive reasons. Differences in consultation rates between men and women were largely eradicated when comparing men and women with similar underlying morbidities.

## Article Summary

### Article focus

- To examine gender differences in consulting in primary care in the UK population using routinely collected general practice data.
- To explore whether gender differences remain when consultation for reproductive reasons or common underlying morbidities are accounted for
- To contribute current evidence on gender difference in using health service

### Key messages

- On average, men consulted less than women between the ages of 16 and 60 years, but not at younger and older ages.
- After controlling for common underlying morbidities (depression and cardiovascular disease), gender differences in the use of primary health care services reduce substantially and are modest.

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**Strengths and limitations of this study**

- Strengths of this study include the use of a UK wide primary care database and the very large study population which is representative of the UK population, enabling a national picture of consultation pattern.
- Limitations include the limited completeness of data recording on variables which may influence gender differences in consultation patterns, such as ethnicity and employment status. Using GP recorded data, it is likely that regular users are over-represented in the analysis.

For peer review only

## Introduction

Average life expectancy is shorter for men than for women in almost all countries, but the magnitude of this female advantage varies geographically and historically, highlighting the importance of understanding what causes gender differences in health. Although biological factors<sup>1-2</sup> and health behaviours<sup>3-4</sup> may provide a partial explanation, it is very widely assumed that men and women have a different propensity to consult<sup>5</sup> and that this may be an important contributor to the gender gap in mortality.<sup>6</sup>

Many large scale studies based on survey data have reported greater use of primary health care services in women.<sup>7-11</sup> In the UK, for example, women aged 16-44 years are twice as likely as men of the same age to have visited their general practitioner (GP) in the previous 12 months.<sup>12</sup> This widely reported “female excess” in consulting has led to an assumption that women are more willing to utilise health services *in all circumstances* and at all ages. However, existing evidence on gender and health service utilisation is mixed,<sup>13-17</sup> and evidence comparing consultation patterns in men and women with similar morbidity is surprisingly sparse and weak.<sup>18-19</sup> Qualitative research has indeed identified that men commonly express a reluctance to consult,<sup>20-21</sup> reinforcing a presumption that men may present with serious disease at a later (and less treatable) stage. Although very few gender comparative studies have been done in this area, there is evidence first that both men *and* women with high levels of multiple morbidity express reluctance to consult or over-use primary care resources,<sup>22</sup> and secondly, that some groups of men express a *readiness* to consult.<sup>23-24</sup> As populations live longer with increasing health service needs, it is important to understand whether there are population subgroups who ‘under-’ or ‘over-consult’ to ensure the most effective use of primary care resources.

In the UK, general practice is usually the first point of access to formal health services, and about 90% of all NHS contacts take place in general practice, with nearly 300 million consultations a year.<sup>25</sup> There has been an upward trend in GP consultation rates between 1995 and 2008, with higher rates recorded for women than men, except in those under 15 years and over 80 years.<sup>7</sup> A similar pattern has been reported in other countries.<sup>26-28</sup> This may reflect greater medicalisation of women’s lives during the reproductive years, gender differences in underlying morbidity, or a different propensity to consult about some kinds of symptoms.<sup>15-16 29</sup> Despite the apparently consistent evidence for higher use of primary health care services in women, two important limitations to the evidence base remain. First, studies rely largely on self-reported survey data, and secondly there is a lack of attention to underlying morbidity. Indeed research evidence on gender and utilisation of health services which use routine data sources is surprisingly scant in the UK and elsewhere.

The aim of this study is to use routinely collected primary care data from the Health Improvement Network (THIN) to explore contemporary consultation patterns in men and women in the UK. We examine whether gender differences in consultation rates are constant across the life-course and across populations living in more or less affluent areas, and whether these differences remain after accounting for consultations related to reproduction and for two common underlying morbidities (depression and cardiovascular disease).

Methods

The Health Improvement Network (THIN) is one of the largest primary care data sources, consisting of electronic records of over ten million patients from more than 500 practices in the UK. Participating practices are broadly representative of UK general practice for patients’ demographic characteristics.<sup>30-31</sup> THIN contains anonymised patient data directly extracted from practices using the Vision general practice system. It is a clinical database which includes every consultation between a health professional and a patient. Data are held on an individual’s age, gender, registration details, clinical symptoms, medical diagnoses, laboratory tests, referrals and prescriptions. The level of consultation recording in THIN is comparable to that from other UK national data sources.<sup>32</sup> THIN also includes an area based deprivation index as a measure of patients’ socioeconomic status. Each postal enumeration district (about 150 households) is assigned a Townsend deprivation score. These areas are divided into national quintiles and patients in THIN are assigned a quintile score.

To be included in the study participating practices had to meet specific criteria, in relation to the completeness of data recording. The first criterion was that participating practices had acceptable computer use:<sup>33</sup> a practice on average continuously records at least one medical record, one additional health data record, and at least two prescriptions for each patient in each year. Secondly, practices’ reporting on mortality rates must be consistent with their patients’ demographic profile.<sup>34</sup> Patients within eligible practices entered the study on the latest of three possible dates: the date of registering with the practice; the date when the practice provided acceptable mortality rates; or the date when the practice attained an acceptable level of data recording. Patients no longer contributed data from the date they were transferred out of the practice, or were recorded as having died, or when the practice has its last data collection.

For this study we identified all direct contacts between clinicians and patients in primary care using Read codes, a hierarchical classification system that includes codes for signs and symptoms, diagnosis, procedures and investigations.<sup>35</sup> A total of 83,722 Read codes were included; these excluded 21,138 codes which concerned records for patient and practice administration, provision of services, hospital procedures and operations. Consultations for reproductive reasons including consulting records for pregnancy, childbirth and contraception were also identified. To further examine the effect of common underlying morbidities on consultation rates amongst men and women, we identified patients who were currently receiving antidepressant prescriptions or treatment for cardiovascular diseases. For each condition we considered patients to be undergoing treatment if they had received two or more relevant prescriptions.<sup>36</sup> We chose to use prescription data rather than Read codes to identify patient with medical diagnoses of these conditions as we are aware that diagnoses are not consistently recorded by general practitioners. We used a cut-off of two or more prescriptions as this group of people were more likely to be taking their medication rather than those who had only ever received one prescription.

## Statistical Analysis

We calculated the annual consultation rate in 2010, using number of consultations recorded and the number of person years as the denominator. If a patient had more than one consultation within a day (whether face-to-face, over the telephone or home visits), we counted only one consultation for that day. We compared rates by gender, age groups and deprivation quintiles. We also calculated consultation rates which excluded all consultations for reproductive reasons (as noted earlier). In exploratory analyses (data not shown) we examined consultation rates for each of three years from 2008-2010; however the patterns were very similar and we therefore focus here on consultations in the most recent year available, 2010.

In order to examine the effect on gender differences in consultation rates of taking account of common underlying morbidities, at an individual patient level we developed generalised Poisson mixed models, including patient age as a nonlinear 5 knot restricted cubic spline, social deprivation (quintiles of Townsend scores) and gender. Practices were included as generalised random effects intercept terms. The log of the number of days followed in 2010 was included as an offset term. The model used a log link and Poisson / Gaussian error terms. We identified those who had at least one reproductive related consultation in 2010 and examined the effect of inclusion of that term or exclusion of those subjects from our models. Further we identified patients who had at least two prescriptions of drugs for CVD or depression.

Analyses were conducted in Stata 12 and SAS 9.2 X64, (SAS Institute, Cary, NC).

## Results

### Consultation rate by gender, age groups and deprivation quintiles

In 2010, there were 1,868,149 men and 1,916,898 women registered with 446 eligible practices. Practice size varied widely ranging from 732 to 29,779 (average practice size was 8487). The crude consultation rate was 32% less in men than in women: men consulted 3152 times per 1000 person years while women consulted 4607 times per 1000 person years (table 1).

The gender difference in consultation rates varied across the life course. As expected, we saw the largest difference in male and female consultation rates in the reproductive years, with a ratio of male to female consultation rate of 0.40 (95% CI 0.392-0.404) between the ages of 21 and 39 years (table 1). However, gender differences in primary care consultation rates were much narrower amongst the youngest (under the age of 21 years, rate ratio= 0.77, 95% CI 0.760-0.780) and the oldest (over the age of 58 years, rate ratio=0.92, 95% CI 0.915-0.927). Indeed, consultation rates in men and women in the oldest age group were quite similar, respectively 6308 per 1000 person years in men and 6851 per 1000 person years in women.

Overall, consultation rates were higher amongst people living in the most deprived areas (5<sup>th</sup> quintile (3946 consultations per 1000 person years) than amongst those in the most affluent areas (1<sup>st</sup> quintile 3806 consultations per person years). However, there was a significant interaction between gender and deprivation status (accounting for gender as spline, and practices as random effects,

p<0.001); the gradient in consultation rates across deprivation quintiles was apparent in women, but not men.

**The effect of reproductive consultations and common underlying morbidities on gender differences in consultation rates**

Many more women (n=239,594) than men (n=829) had consulted for reproductive reasons, and around twice as many were receiving antidepressant medication (women=173,407; men=76,602) (Table 2); the numbers of men and women who were receiving cardiovascular medication were more similar (Table 2). Gender differences in consulting rates amongst these three groups of patients were much smaller than those seen in the general population (table 1). For example, amongst people with depression, the consultation rate was 9102 per 1000 person year in men and 9961 per 1000 person year in women. The differences in consultation rates between men and women were also reduced when reproductive related consultations were excluded, although a considerable gap between male and female consultation rates remained between the ages of 15 and 60 years (figure 1).

Overall, after conditioning for age and deprivation quintile, there remained a substantial difference in consultation between men and women (relative risk (RR) 0.719 (95% CI 0.718 to 0.720, table 3)). This was in part ameliorated when reproductive consultations were accounted for (RR=0.81, 95%CI 0.809 to 0.811). When we further accounted for common underlying morbidities, the relative risks were much closer to unity: thus after accounting for depression the RR was 0.916 (95% CI 0.913 to 0.918) and after accounting for cardiovascular disease the RR was 0.950 (95%CI 0.948 to 0.952) (table 3). In addition, gender differences in consulting became even smaller when reproductive consultations were also accounted for amongst those on antidepressant medications (RR=0.951, 95% CI 0.948-0.954); additional adjustment for consultations for reproductive reasons had a little impact on gender differences in consulting for those with CVD.

**Discussion**

This study examined the extent of gender differences in the use of primary care services in the UK over the life-course and amongst people from different socio-economic backgrounds, before and after taking account of consultations for reproductive reasons and two common underlying morbidities, cardiovascular diseases and depression). Overall, as expected, we found that men's consultation rates were over 30% lower than women's, confirming the gender pattern in primary care consultations reported in earlier self-report surveys.<sup>11</sup> The magnitude of gender differences in primary care consulting varied by age; there was very little difference in childhood and older age, and much higher rates of consulting in women than men during the reproductive and mid-life years. The variation in gender differences over the life-course which we observed is very similar to that reported in another study using an alternative source of routinely collected data on primary care consultations in the UK.<sup>7</sup> It is often suggested that women's higher rates of health service contacts can be attributed to consulting for reproductive health,<sup>37</sup> but in these data consultations for reproductive reasons only partially explained the large gap in consulting between men and women in mid-life, reflecting findings from small-scale studies using self-reported data.<sup>5 38</sup> The gender



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3 difference in consulting also varied by deprivation status, reflecting a socio-economic gradient in  
4 consulting rates amongst women but not men.  
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6 Another important finding of the study is that men and women with similar underlying morbidities  
7 (depression and cardiovascular disease) differ much less in their use of primary health care services  
8 than men and women in the population as a whole. For this analysis, we choose two common  
9 conditions, which are usually managed by general practitioners, both of which differ by gender for  
10 morbidity and mortality. Compared to an overall male to female consultation rate ratio of 0.72, the  
11 ratio was 0.92 for patients with depression and 0.95 for patients with cardiovascular disease. This  
12 mirrors findings from smaller studies which have relied on self-reported data on morbidity and  
13 consulting which suggest that gender differences in consulting are small after taking account of  
14 underlying morbidity.<sup>19 39</sup> Our findings suggest that some of the “excess” in female consulting may  
15 reflect greater levels of ‘need’ in relation to depression in women. However, given that we identified  
16 patients with depression through their medication records, it is important to note that previous  
17 studies have reported that women are more likely to receive a prescription when consulting their  
18 GPs.<sup>40</sup> Well-designed experimental studies in the UK and the USA have shown how GPs take account  
19 of gender, ethnicity and age in assessing the likely importance of symptom presentations: given the  
20 same presentations of symptoms of coronary heart disease (CHD), GPs were more likely to attribute  
21 these to CHD in men and to have a higher level of certainty about their diagnosis.<sup>41-43</sup> Although  
22 gender comparative evidence on consulting for other potentially fatal diseases is sparse,<sup>44</sup> a US study  
23 which compared the consulting histories of men and women diagnosed with colorectal cancer found  
24 that on average women delayed *longer* than men after first noticing their symptoms, and made  
25 more visits to the doctor before gaining their diagnosis. A more recent UK study of a consecutive  
26 series of lung cancer patients at three Scottish hospitals found that gender was not a predictor of  
27 time from first noticing symptoms to consulting a general practitioner.<sup>45</sup>  
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34 The strengths of this study include the use of a UK wide primary care database and the large study  
35 population which is representative of the UK population, enabling a national picture of consultation  
36 pattern. However, there are also limitations which should be noted. The aim of this study was to  
37 explore the effect of gender on the use of primary care services after considering other confounding  
38 characteristics. We were not able to examine to what extent gender differences in consultation  
39 pattern are influenced by factors such as ethnicity and employment status because of limited  
40 completeness of data recording on these variables. Studies have shown the effect of ethnicity and  
41 gender concordance (similarity in gender of doctor and patients) on the use of services.<sup>46-47</sup> Secondly,  
42 our analyses focused on the utilisation of primary care health services, but were not able to examine  
43 fully the level of clinical needs. Thirdly, THIN includes GP recorded data, which cover consultation  
44 information about patients who have been in contact with GPs or nurses, and it is likely that regular  
45 users are over-represented in the analysis. Finally, as noted earlier, we included patients who had 2  
46 or more prescriptions as a proxy for underlying morbidity, but this may have failed to identify  
47 patients with milder depression/ cardiovascular disease if their condition was not severe enough to  
48 warrant medication.  
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53 Despite these limitations, this study provides much needed information on recent primary care use  
54 in men and women. It shows that gender differences in consulting are not universal; indeed the  
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magnitude of gender differences are modest at some stages of the life course, and when account is taken of underlying morbidity.

**Conclusions**

Differences in consultations between men and women are most marked between the ages of 16 and 60 years, and these differences are only partially accounted for when consultations for reproductive health are considered. However, gender differences in consultations rates in patients with cardiovascular disease and depression are relatively small. General practitioners need to be aware in planning their delivery of health care that the gender difference in primary care health service utilisation are not constant and do not simply reflect a greater and universal propensity for women to consult more readily than men.

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**Competing interests** None

**Ethics approval** This study received approval from EPIC Scientific Review Committee.

**Data sharing statement** No additional data are available

**Contributorship statement:** KH originally conceived the study. YW extracted the data from the Health Improvement Network database and undertook the initial data analyses. NF conducted statistical analyses using Poisson mixed models. All authors contributed to the study design and decisions on the interpretation of results. YW HK contributed to the drafting of the manuscript. All authors contributed to the critical revision of the manuscript and approved the final version of the manuscript prior to submission. KH is the guarantor.

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**Table 1:** Crude consultation rate per 1000 person year in 2010, by gender, age groups and deprivation

	Men					Women					M/F ratio (95% CI)
	No.	/	person year	Rate	(95% CI)	No.	/	person year	Rate	(95% CI)	
Overall	5361100		1700883	<b>3152</b>	(3149 3155)	8001121		1736618	<b>4607</b>	(4604 4610)	<b>0.68</b> (0.68 0.69)
Age quartiles											
0-20	833499		415772.6	<b>2005</b>	(2000 2009)	1033997		396957.6	<b>2605</b>	(2600 2610)	<b>0.77</b> (0.76 0.78)
21-39	705785		425352.9	<b>1659</b>	(1655 1663)	1766972		423892.6	<b>4168</b>	(4162 4174)	<b>0.40</b> (0.39 0.40)
40-57	1225957		448215.3	<b>2735</b>	(2730 2740)	1914647		436238.1	<b>4389</b>	(4383 4395)	<b>0.62</b> (0.62 0.63)
58+	2595859		411542.2	<b>6308</b>	(6300 6315)	3285505		479529.5	<b>6851</b>	(6844 6859)	<b>0.92</b> (0.91 0.93)
Deprivation quintiles <sup>1</sup>											
1	1380470		432255.4	<b>3194</b>	(3188 3199)	1940611		440269.4	<b>4408</b>	(4401 4414)	<b>0.72</b> (0.72 0.73)
2	1176355		363717.9	<b>3234</b>	(3228 3240)	1699076		373145.0	<b>4553</b>	(4546 4560)	<b>0.71</b> (0.70 0.72)
3	1073696		343302.0	<b>3127</b>	(3122 3133)	1633282		351722.6	<b>4644</b>	(4636 4651)	<b>0.67</b> (0.67 0.68)
4	940676		302437.5	<b>3110</b>	(3104 3117)	1478517		308806.1	<b>4788</b>	(4780 4795)	<b>0.65</b> (0.64 0.66)
5	655705		212989.0	<b>3078</b>	(3071 3086)	1018719		211267.7	<b>4822</b>	(4812 4831)	<b>0.64</b> (0.63 0.65)

1. 114,537 patients' socio-deprivation data were missing.

**Table 2:** Number of patients and crude consultation rates amongst 3 groups of patients in 2010

	Men	Women	All
<b>No. of patients who had at least one consultation for reproductive reasons (%)</b>	829 (0.04)	239,594 (12.50)	240,423(6.35)
age (median years)	37.5	29.5	29.5
crude consultation rate (95% CI)	5999 (5832 6169)	6283 (6272 6293)	6282 (6272 6292)
<b>No. of patients with CVD (%)</b>	221,734 (11.87)	254,831 (13.29)	476,565 (12.59)
age (median years)	66.5	68.5	67.5
crude consultation rate (95%CI)	9441 (9429 9454)	10180 (10167 10192)	9836 (9827 9845)
<b>No. of patients with depression (%)</b>	76,602 (4.10)	173,407 (9.05)	250,009 (6.61)
age (median years)	51.5	50.5	50.5
crude consultation rate (95%CI)	9102 (9081 9124)	9961 (9946 9976)	9698 (9686 9711)

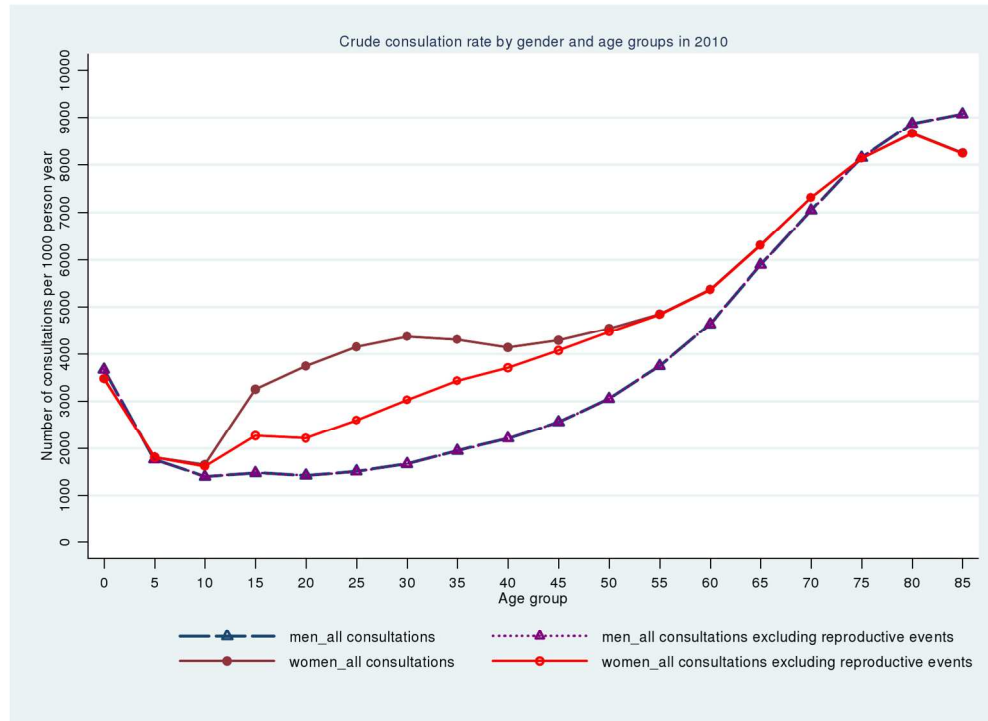
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**Table 3:** Gender differences in consultations adjusted for age and deprivation

	Relative Risk (95 % CI)	No. Males (%)	No. Females (%)
All patients	0.719 (0.718 0.720)	1868149 (49.36)	1916898 (50.64)
Patients conditioning for consulting for reproductive reasons	0.810 (0.809 0.811)	1868149 (49.36)	1916898 (50.64)
Patients with CVD	0.950 (0.948 0.952)	221734 (46.53)	254831 (53.47)
Patients with depression	0.916 (0.913 0.918)	76602 (30.64)	173407 (69.36)
Patients with CVD and adjusted for reproductive reasons	0.957 (0.955 0.959)	221734 (46.53)	254831 (53.47)
Patients with depression and adjusted for reproductive reasons	0.950 (0.947 0.953)	76602 (30.64)	173407 (69.36)



**Figure 1:** Consultation rate per person year by gender and age (5 years age band) in 2010



140x112mm (300 x 300 DPI)

STROBE Statement—Checklist of items that should be included in reports of *cross-sectional studies*

	Item No	Recommendation
<b>Title and abstract</b>	1	(a) Indicate the study’s design with a commonly used term in the title or the abstract (b) Provide in the abstract an informative and balanced summary of what was done and what was found
<b>Introduction</b>		
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported
Objectives	3	State specific objectives, including any prespecified hypotheses
<b>Methods</b>		
Study design	4	Present key elements of study design early in the paper
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection
Participants	6	(a) Give the eligibility criteria, and the sources and methods of selection of participants
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable
Data sources/measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group
Bias	9	Describe any efforts to address potential sources of bias
Study size	10	Explain how the study size was arrived at
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why
Statistical methods	12	(a) Describe all statistical methods, including those used to control for confounding (b) Describe any methods used to examine subgroups and interactions (c) Explain how missing data were addressed (d) If applicable, describe analytical methods taking account of sampling strategy (e) Describe any sensitivity analyses
<b>Results</b>		
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed (b) Give reasons for non-participation at each stage (c) Consider use of a flow diagram
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders (b) Indicate number of participants with missing data for each variable of interest
Outcome data	15*	Report numbers of outcome events or summary measures
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included (b) Report category boundaries when continuous variables were categorized (c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses

<b>Discussion</b>		
Key results	18	Summarise key results with reference to study objectives
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence
Generalisability	21	Discuss the generalisability (external validity) of the study results
<b>Other information</b>		
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based

\*Give information separately for exposed and unexposed groups.

**Note:** An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at <http://www.plosmedicine.org/>, Annals of Internal Medicine at <http://www.annals.org/>, and Epidemiology at <http://www.epidem.com/>). Information on the STROBE Initiative is available at [www.strobe-statement.org](http://www.strobe-statement.org).



**Do men consult less than women? An analysis of routinely collected UK general practice data**

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**Title Page**

**Do men consult less than women? An analysis of routinely collected UK general practice data**

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## Abstract

**Objective:** To examine whether gender differences in primary care consultation rates a) vary by age and deprivation status, and b) diminish when consultation for reproductive reasons or common underlying morbidities are accounted for.

**Design:** Cross sectional study of a cohort of patients registered with general practice.

**Setting:** UK primary care

**Subjects:** Patients (1,869,149 men, 1,916,898 women) registered with 446 eligible practices in 2010

**Primary outcome measures:** Primary care consultation rate

**Results:** This study analyses routinely collected primary care consultation data. The crude consultation rate was 32% lower in men than women. The magnitude of gender difference varied across the life-course, and there was no "excess" female consulting in early and later life. The greatest gender gap in primary care consultations was seen amongst those aged between 16 and 60 years. Gender differences in consulting were higher in people from more deprived areas than amongst those from more affluent areas. Accounting for reproductive related consultations diminished but did not eradicate the gender gap. However, consultation rates in men and women who had comparable underlying morbidities (as assessed by receipt of medication) were similar; men in receipt of antidepressant medication were only 8% less likely to consult than women in receipt of antidepressant medication (relative risk 0.916, 95% confidence interval (CI) 0.913-0.918), and men with in receipt of medication to treat cardiovascular disease were just 5% less likely to consult (RR=0.950, 95%CI 0.948-0.952) than women receiving similar medication. These small gender differences diminished further, particularly for depression (RR=0.950, 95% CI 0.947-0.953), after also taking account of reproductive consultations.

**Conclusions:** Overall gender differences in consulting are most marked between the ages of 16 and 60 years; these differences are only partially accounted for by consultations for reproductive reasons. Differences in consultation rates between men and women were largely eradicated when comparing men and women with in receipt of medication for similar underlying morbidities.

## Article Summary

### Article focus

- To examine gender differences in consulting in primary care in the UK population using routinely collected general practice data.
- To explore whether gender differences remain when consultation for reproductive reasons or treatment for common underlying morbidities are accounted for
- To contribute current evidence on gender difference in health service use

### Key messages

- On average, men consulted less than women between the ages of 16 and 60 years, but not at younger and older ages.

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- After controlling for medication for common underlying morbidities (depression and cardiovascular disease), gender differences in the use of primary health care services reduce substantially and are modest.

**Strengths and limitations of this study**

- Strengths of this study include the use of a UK wide primary care database and the very large study population which is representative of the UK population, enabling a national picture of consultation pattern.
- Limitations include the limited completeness of data recording on variables which may influence gender differences in consultation patterns, such as ethnicity and employment status. Using GP recorded data, it is likely that regular users are over-represented in the analysis.

For peer review only

## Introduction

Average life expectancy is shorter for men than for women in almost all countries, but the magnitude of this female advantage varies geographically and historically, highlighting the importance of understanding what causes gender differences in health. Although biological factors<sup>1-2</sup> and health behaviours<sup>3-4</sup> may provide a partial explanation, it is very widely assumed that men and women have a different propensity to consult<sup>5</sup> and that this may be an important contributor to the gender gap in mortality.<sup>6</sup>

Many large scale studies based on survey data have reported greater use of primary health care services in women.<sup>7-11</sup> In the UK, for example, women aged 16-44 years are twice as likely as men of the same age to have visited their general practitioner (GP) in the previous 12 months.<sup>12</sup> This widely reported “female excess” in consulting has led to an assumption that women are more willing to utilise health services *in all circumstances* and at all ages. However, existing evidence on gender and health service utilisation is mixed,<sup>13-17</sup> and evidence comparing consultation patterns in men and women with similar morbidity is surprisingly sparse and weak.<sup>18-19</sup> Qualitative research has indeed identified that men commonly express a reluctance to consult,<sup>20-21</sup> reinforcing a presumption that men may present with serious disease at a later (and less treatable) stage. Although very few gender comparative studies have been done in this area, there is evidence first that both men *and* women with high levels of multiple morbidity express reluctance to consult or over-use primary care resources, in the UK at least,<sup>22</sup> and secondly, that some groups of men express a *readiness* to consult.<sup>23-24</sup> As populations live longer with increasing health service needs, it is important to understand whether there are population subgroups who ‘under-’ or ‘over-consult’ to ensure the most effective use of primary care resources.

In the UK, general practice is usually the first point of access to formal health services, and about 90% of all NHS contacts take place in general practice, with nearly 300 million consultations a year.<sup>25</sup> There has been an upward trend in GP consultation rates between 1995 and 2008, with higher rates recorded for women than men, except in those under 15 years and over 80 years.<sup>7</sup> A similar pattern has been reported in other countries.<sup>26-28</sup> This may reflect greater medicalisation of women’s lives during the reproductive years, gender differences in underlying morbidity, or a different propensity to consult about some kinds of symptoms.<sup>15-16 29</sup> Despite the apparently consistent evidence for higher use of primary health care services in women, two important limitations to the evidence base remain. First, studies rely largely on self-reported survey data, and secondly there is a lack of attention to underlying morbidity. Indeed research evidence on gender and utilisation of health services which use routine data sources is surprisingly scant in the UK and elsewhere.

The aim of this study is to use routinely collected primary care data from the Health Improvement Network (THIN) to explore contemporary consultation patterns in men and women in the UK. We examine whether gender differences in consultation rates are constant across the life-course and across populations living in more or less affluent areas, and whether these differences remain after accounting for consultations related to reproduction and for two common underlying morbidities (depression and cardiovascular disease).



Methods

The Health Improvement Network (THIN) is one of the largest primary care data sources, consisting of electronic records of over ten million patients from more than 500 practices in the UK. Participating practices are broadly representative of UK general practice for patients’ demographic characteristics.<sup>30-31</sup> THIN contains anonymised patient data directly extracted from practices using the Vision general practice system. It is a clinical database which includes every consultation between a health professional and a patient. Data are held on an individual’s age, gender, registration details, clinical symptoms, medical diagnoses, laboratory tests, referrals and prescriptions. The level of consultation recording in THIN is comparable to that from other UK national data sources.<sup>32</sup> THIN also includes an area based deprivation index as a measure of patients’ socioeconomic status. Each postal enumeration district (about 150 households) is assigned a Townsend deprivation score. These areas are divided into national quintiles and patients in THIN are assigned a quintile score.

To be included in the study participating practices had to meet specific criteria, in relation to the completeness of data recording. The first criterion was that participating practices had acceptable computer use:<sup>33</sup> a practice on average continuously records at least one medical record, one additional health data record, and at least two prescriptions for each patient in each year. Secondly, practices’ reporting on mortality rates must be consistent with their patients’ demographic profile.<sup>34</sup> Patients within eligible practices entered the study on the latest of three possible dates: the date of registering with the practice; the date when the practice provided acceptable mortality rates; or the date when the practice attained an acceptable level of data recording. Patients no longer contributed data from the date they were transferred out of the practice, or were recorded as having died, or when the practice has its last data collection.

For this study we identified all direct contacts between clinicians and patients in primary care using Read codes, a hierarchical classification system that includes codes for signs and symptoms, diagnosis, procedures and investigations.<sup>35</sup> A total of 83,722 Read codes were included; these excluded 21,138 codes which concerned records for patient and practice administration, provision of services, hospital procedures and operations. Consultations for reproductive reasons were also identified: these included all consultations related to normal and abnormal pregnancy, childbirth and post-natal consultations (including complications following childbirth) and contraception) . To further examine the effect of underlying morbidities on consultation rates amongst men and women, we identified patients who were in receipt of medication to treat depression and cardiovascular disease as a clinical marker for patients with depression and CVD. These two conditions were selected as morbidities which: a) occur in both men and women, but there are gender differences in morbidity and mortality related to both conditions; b) are common in adult life; and c) are treated within general practice. Hence, for each condition we considered patients to be actively undergoing treatment for depression or cardiovascular if they had received two or more relevant prescriptions.<sup>36</sup> We chose to use prescription data rather than Read codes to identify patient with medical diagnoses of these conditions as we are aware that diagnoses are not consistently recorded by general practitioners; whilst some doctors record a diagnosis such as depression each time a patient consults, others will not include the diagnosis on the patient record for a consultation if it has been previously recorded, whereas medications are recorded more consistently. Furthermore, the issue of a

prescription suggests that the underlying morbidity is sufficiently serious to warrant medication. We used a cut-off of two or more prescriptions as this group of people were more likely to be taking their medication rather than those who had only ever received one prescription.

## Statistical Analysis

We calculated the annual consultation rate in 2010, using number of consultations recorded and the number of person years as the denominator. If a patient had more than one consultation within a day (whether face-to-face, over the telephone or home visits), we counted only one consultation for that day. We compared rates by gender, age groups and deprivation quintiles. We also calculated consultation rates which excluded all consultations for reproductive reasons (as noted earlier). In exploratory analyses (data not shown) we examined consultation rates for each of three years from 2008-2010; however the patterns were very similar and we therefore focus here on consultations in the most recent year available, 2010.

In order to examine the effect on gender differences in consultation rates of taking account of common underlying morbidities, as described above, at an individual patient level we developed generalised Poisson mixed models, including patient age as a nonlinear 5 knot restricted cubic spline, social deprivation (quintiles of Townsend scores) and gender. Practices were included as generalised random effects intercept terms. The log of the number of days followed in 2010 was included as an offset term. The model used a log link and Poisson / Gaussian error terms. We identified those who had at least one reproductive related consultation in 2010 and examined the effect of inclusion of that term or exclusion of those subjects from our models. Further we identified patients who had at least two prescriptions of drugs for CVD or depression.

Analyses were conducted in Stata 12 and SAS 9.2 X64, (SAS Institute, Cary, NC).

## Results

### Consultation rate by gender, age groups and deprivation quintiles

In 2010, there were 1,868,149 men and 1,916,898 women registered with 446 eligible practices. Practice size varied widely ranging from 732 to 29,779 (average practice size was 8487). The crude consultation rate was 32% less in men than in women: men consulted 3152 times per 1000 person years while women consulted 4607 times per 1000 person years (table 1).

The gender difference in consultation rates varied across the life course. As expected, we saw the largest difference in male and female consultation rates in the reproductive years, with a ratio of male to female consultation rate of 0.40 (95% CI 0.392-0.404) between the ages of 21 and 39 years (table 1). However, gender differences in primary care consultation rates were much narrower amongst the youngest (under the age of 21 years, rate ratio= 0.77, 95% CI 0.760-0.780) and the oldest (over the age of 58 years, rate ratio=0.92, 95% CI 0.915-0.927). Indeed, consultation rates in

men and women in the oldest age group were quite similar, respectively 6308 per 1000 person years in men and 6851 per 1000 person years in women.

Overall, consultation rates were higher amongst people living in the most deprived areas (5<sup>th</sup> quintile (3946 consultations per 1000 person years) than amongst those in the most affluent areas (1<sup>st</sup> quintile 3806 consultations per person years). However, there was a significant interaction between gender and deprivation status (accounting for gender as spline, and practices as random effects,  $p<0.001$ ); the gradient in consultation rates across deprivation quintiles was apparent in women, but not men.

**The effect of reproductive consultations and common underlying morbidities on gender differences in consultation rates**

Many more women (n=239,594) than men (n=829) had consulted for reproductive reasons, and around twice as many were receiving antidepressant medication (women=173,407; men=76,602) (Table 2); the numbers of men and women who were receiving cardiovascular medication were more similar (Table 2). Gender differences in consulting rates amongst these three groups of patients were much smaller than those seen in the general population (table 1). For example, amongst people in receipt of medication for depression , the consultation rate was 9102 per 1000 person year in men and 9961 per 1000 person year in women. The differences in consultation rates between men and women were also reduced when reproductive related consultations were excluded, although a considerable gap between male and female consultation rates remained between the ages of 15 and 60 years (figure 1).

Overall, after conditioning for age and deprivation quintile, there remained a substantial difference in consultation between men and women (relative risk (RR) 0.719 (95% CI 0.718 to 0.720, table 3)). This was in part ameliorated when reproductive consultations were accounted for (RR=0.81, 95%CI 0.809 to 0.811). When we further accounted for common underlying morbidities, the relative risks were much closer to unity: thus after accounting for being in receipt of medication for depression the RR was 0.916 (95% CI 0.913 to 0.918) and after accounting for being in receipt of medication for cardiovascular disease the RR was 0.950 (95%CI 0.948 to 0.952) (table 3). In addition, gender differences in consulting became even smaller when reproductive consultations were also accounted for amongst those on antidepressant medications (RR=0.951, 95% CI 0.948-0.954); additional adjustment for consultations for reproductive reasons had a little impact on gender differences in consulting for those with CVD.

**Discussion**

This study examined the extent of gender differences in the use of primary care services in the UK over the life-course and amongst people from different socio-economic backgrounds, before and after taking account of consultations for reproductive reasons and two common underlying morbidities, cardiovascular diseases and depression, as represented by being in receipt of medication for these conditions). Overall, as expected, we found that men’s consultation rates were over 30% lower than women’s, confirming the gender pattern in primary care consultations reported

in earlier self-report surveys.<sup>11</sup> The magnitude of gender differences in primary care consulting varied by age; there was very little difference in childhood and older age, and much higher rates of consulting in women than men during the reproductive and mid-life years. The variation in gender differences over the life-course which we observed is very similar to that reported in another study using an alternative source of routinely collected data on primary care consultations in the UK.<sup>7</sup> These studies together provide strong evidence that, on average during the working years men have fewer contacts with their GPs than women do. It is often suggested that women's higher rates of health service contacts can be attributed to consulting for reproductive health,<sup>37</sup> but in these data consultations for reproductive reasons only partially explained the large gap in consulting between men and women in mid-life, reflecting findings from small-scale studies using self-reported data.<sup>5 38</sup> The gender difference in consulting also varied by deprivation status, reflecting a socio-economic gradient in consulting rates amongst women but not men. This finding was unexpected and warrants further exploration.

Another important finding of the study is that men and women with similar underlying morbidities (depression and cardiovascular disease) differ much less in their use of primary health care services than men and women in the population as a whole. For this analysis, we choose two common conditions, which are usually managed by general practitioners, both of which differ by gender for morbidity and mortality. Compared to an overall male to female consultation rate ratio of 0.72, the ratio was 0.92 for patients in receipt of medication for depression and 0.95 for patients in receipt of medication for cardiovascular disease. This mirrors findings from smaller studies which have relied on self-reported data on morbidity and consulting which suggest that gender differences in consulting are small after taking account of underlying morbidity.<sup>19 39</sup> Our findings suggest that some of the "excess" in female consulting may reflect greater levels of 'need' in relation to depression in women. However, given that we identified patients with depression through their medication records, it is important to note that previous studies have reported that women are more likely to receive a prescription when consulting their GPs.<sup>40</sup> Well-designed experimental studies in the UK and the USA have shown how GPs take account of gender, ethnicity and age in assessing the likely importance of symptom presentations: given the same presentations of symptoms of coronary heart disease (CHD), GPs were more likely to attribute these to CHD in men and to have a higher level of certainty about their diagnosis.<sup>41-43</sup> Although gender comparative evidence on consulting for other potentially fatal diseases is sparse,<sup>44</sup> a US study which compared the consulting histories of men and women diagnosed with colorectal cancer found that on average women delayed *longer* than men after first noticing their symptoms, and made more visits to the doctor before gaining their diagnosis. A more recent UK study of a consecutive series of lung cancer patients at three Scottish hospitals found that gender was not a predictor of time from first noticing symptoms to consulting a general practitioner.<sup>45</sup>

The strengths of this study include the use of a UK wide primary care database and the large study population which is representative of the UK population, enabling a national picture of consultation pattern. However, there are also limitations which should be noted. The aim of this study was to explore the effect of gender on the use of primary care services after considering other confounding characteristics. We were not able to examine to what extent gender differences in consultation pattern are influenced by factors such as ethnicity and employment status because of limited completeness of data recording on these variables. Studies have shown the effect of ethnicity and

gender concordance (similarity in gender of doctor and patients) on the use of services.<sup>46-47</sup> Secondly, our analyses focused on the utilisation of primary care health services, but were not able to examine fully the level of clinical needs. Thirdly, THIN includes GP recorded data, which cover consultation information about patients who have been in contact with GPs or nurses, and it is likely that regular users are over-represented in the analysis. Finally, as noted earlier, we included patients who had two or more prescriptions as a proxy for underlying morbidity, but this may have failed to identify patients with milder depression/ cardiovascular disease if their condition was not severe enough to warrant medication. In addition, some prescriptions such as antidepressants can be used to treat other conditions, such as chronic pain. Whilst this is likely to occur relatively infrequently, if there are gender differences in these other conditions, this could differentially affect the specificity of controlling for use of that medication in relation to gender differences in consulting for the conditions of interest. Furthermore, if there is any difference in the compliance with medication by gender, with women being more likely to pick up a second prescription, then this could explain some of the diminution in gender differences in consulting when controlling for medication use.

Despite these limitations, this study provides much needed information on recent primary care use in men and women. It shows that gender differences in consulting are not universal; indeed the magnitude of gender differences are modest at some stages of the life course, and when account is taken of underlying morbidity.

**Conclusions**

Differences in consultations between men and women are most marked between the ages of 16 and 60 years, confirming that on average men have fewer contacts with general practitioners in early adulthood and mid-life, a difference that is only partially accounted for when consultations for reproductive health are considered. However, gender differences in consultations rates in patients in receipt of medication for cardiovascular disease and depression are relatively small, suggesting men and women with common morbidities may have more similar patterns of consulting. General practitioners need to be aware in planning their delivery of health care that the gender difference in primary care health service utilisation are not constant and do not simply reflect a greater and universal propensity for women to consult more readily than men.

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**Data sharing statement** No additional data are available

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**Table 1:** Crude consultation rate per 1000 person year in 2010, by gender, age groups and deprivation

	Men					Women					M/F ratio (95% CI)
	No.	/	person year	Rate	(95% CI)	No.	/	person year	Rate	(95% CI)	
Overall	5361100		1700883	<b>3152</b>	(3149 3155)	8001121		1736618	<b>4607</b>	(4604 4610)	<b>0.68</b> (0.68 0.69)
Age quartiles											
0-20	833499		415772.6	<b>2005</b>	(2000 2009)	1033997		396957.6	<b>2605</b>	(2600 2610)	<b>0.77</b> (0.76 0.78)
21-39	705785		425352.9	<b>1659</b>	(1655 1663)	1766972		423892.6	<b>4168</b>	(4162 4174)	<b>0.40</b> (0.39 0.40)
40-57	1225957		448215.3	<b>2735</b>	(2730 2740)	1914647		436238.1	<b>4389</b>	(4383 4395)	<b>0.62</b> (0.62 0.63)
58+	2595859		411542.2	<b>6308</b>	(6300 6315)	3285505		479529.5	<b>6851</b>	(6844 6859)	<b>0.92</b> (0.91 0.93)
Deprivation quintiles <sup>1</sup>											
1	1380470		432255.4	<b>3194</b>	(3188 3199)	1940611		440269.4	<b>4408</b>	(4401 4414)	<b>0.72</b> (0.72 0.73)
2	1176355		363717.9	<b>3234</b>	(3228 3240)	1699076		373145.0	<b>4553</b>	(4546 4560)	<b>0.71</b> (0.70 0.72)
3	1073696		343302.0	<b>3127</b>	(3122 3133)	1633282		351722.6	<b>4644</b>	(4636 4651)	<b>0.67</b> (0.67 0.68)
4	940676		302437.5	<b>3110</b>	(3104 3117)	1478517		308806.1	<b>4788</b>	(4780 4795)	<b>0.65</b> (0.64 0.66)
5	655705		212989.0	<b>3078</b>	(3071 3086)	1018719		211267.7	<b>4822</b>	(4812 4831)	<b>0.64</b> (0.63 0.65)

1. 114,537 patients' socio-deprivation data were missing.

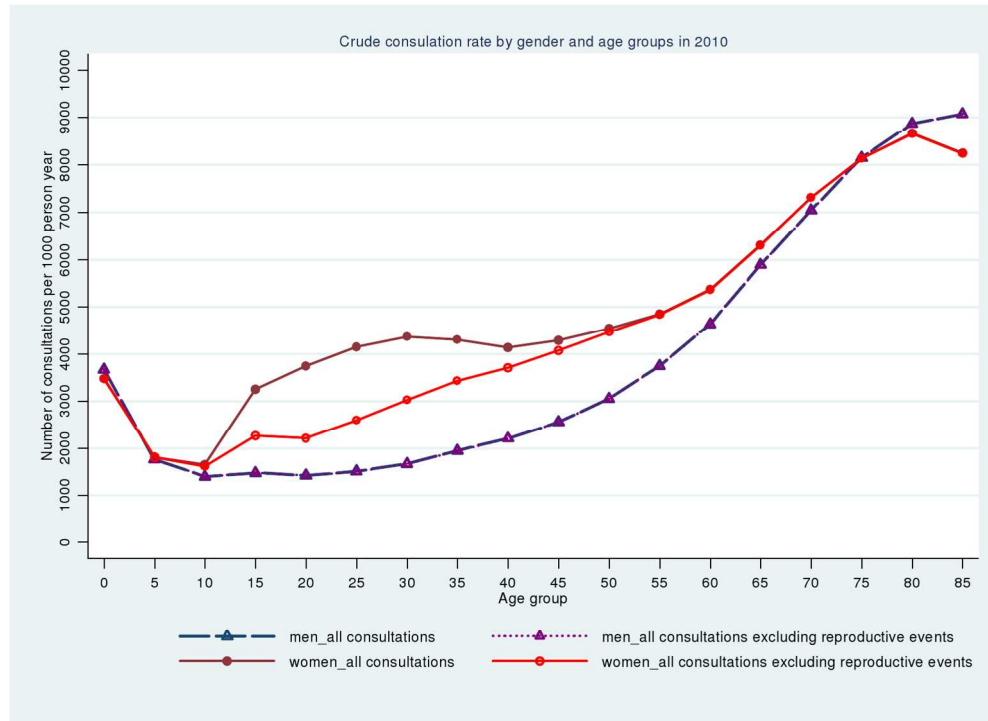
**Table 2:** Number of patients and crude consultation rates amongst 3 groups of patients in 2010

	Men	Women	All
<b>No. of patients who had at least one consultation for reproductive reasons (%)</b>	829 (0.04)	239,594 (12.50)	240,423(6.35)
age (median years)	37.5	29.5	29.5
crude consultation rate (95% CI)	5999 (5832 6169)	6283 (6272 6293)	6282 (6272 6292)
<b>No. of patients in receipt of medication for CVD (%)</b>	221,734 (11.87)	254,831 (13.29)	476,565 (12.59)
age (median years)	66.5	68.5	67.5
crude consultation rate (95%CI)	9441 (9429 9454)	10180 (10167 10192)	9836 (9827 9845)
<b>No. of patients in receipt of medication for depression (%)</b>	76,602 (4.10)	173,407 (9.05)	250,009 (6.61)
age (median years)	51.5	50.5	50.5
crude consultation rate (95%CI)	9102 (9081 9124)	9961 (9946 9976)	9698 (9686 9711)

**Table 3:** Gender differences in consultations adjusted for age and deprivation

	Relative Risk (95 % CI)	No. Males (%)	No. Females (%)
All patients	0.719 (0.718 0.720)	1868149 (49.36)	1916898 (50.64)
Patients conditioning for consulting for reproductive reasons	0.810 (0.809 0.811)	1868149 (49.36)	1916898 (50.64)
Patients in receipt of medication for CVD	0.950 (0.948 0.952)	221734 (46.53)	254831 (53.47)
Patients in receipt of medication for depression	0.916 (0.913 0.918)	76602 (30.64)	173407 (69.36)
Patients in receipt of medication for CVD and adjusted for reproductive reasons	0.957 (0.955 0.959)	221734 (46.53)	254831 (53.47)
Patients in receipt of medication for depression and adjusted for reproductive reasons	0.950 (0.947 0.953)	76602 (30.64)	173407 (69.36)

**Figure 1:** Consultation rate per person year by gender and age (5 years age band) in 2010



140x112mm (300 x 300 DPI)

**Title Page**

**Do men consult less than women? An analysis of routinely collected UK general practice data**

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Keywords: gender, consultation, use of health service, primary care

Word count: 2,770

## Abstract

**Objective:** To examine whether gender differences in primary care consultation rates a) vary by age and deprivation status, and b) diminish when consultation for reproductive reasons or common underlying morbidities are accounted for.

**Design:** Cross sectional study of a cohort of patients registered with general practice.

**Setting:** UK primary care

**Subjects:** Patients (1,869,149 men, 1,916,898 women) registered with 446 eligible practices in 2010

**Primary outcome measures:** Primary care consultation rate

**Results:** This study analyses routinely collected primary care consultation data. The crude consultation rate was 32% lower in men than women. The magnitude of gender difference varied across the life-course, and there was no “excess” female consulting in early and later life. The greatest gender gap in primary care consultations was seen amongst those aged between 16 and 60 years. Gender differences in consulting were higher in people from more deprived areas than amongst those from more affluent areas. Accounting for reproductive related consultations diminished but did not eradicate the gender gap. However, consultation rates in men and women who had comparable underlying morbidities (as assessed by receipt of medication) were similar; men ~~with depression~~ in receipt of antidepressant medication were only 8% less likely to consult than women in receipt of antidepressant medication with depression (relative risk 0.916, 95% confidence interval (CI) 0.913-0.918), and men with in receipt of medication to treat cardiovascular disease were just 5% less likely to consult (RR=0.950, 95%CI 0.948-0.952) than women ~~with cardiovascular disease~~ receiving similar medication. These small gender differences diminished further, particularly for depression (RR=0.950, 95% CI 0.947-0.953), after also taking account of reproductive consultations.

**Conclusions:** Overall gender differences in consulting are most marked between the ages of 16 and 60 years; these differences are only partially accounted for by consultations for reproductive reasons. Differences in consultation rates between men and women were largely eradicated when comparing men and women with in receipt of medication for similar underlying morbidities.

## Article Summary

### Article focus

- To examine gender differences in consulting in primary care in the UK population using routinely collected general practice data.
- To explore whether gender differences remain when consultation for reproductive reasons or treatment for common underlying morbidities are accounted for
- To contribute current evidence on gender difference in ~~using~~ health service use

### Key messages

- On average, men consulted less than women between the ages of 16 and 60 years, but not at younger and older ages.

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- After controlling for medication for common underlying morbidities (depression and cardiovascular disease), gender differences in the use of primary health care services reduce substantially and are modest.

**Strengths and limitations of this study**

- Strengths of this study include the use of a UK wide primary care database and the very large study population which is representative of the UK population, enabling a national picture of consultation pattern.
- Limitations include the limited completeness of data recording on variables which may influence gender differences in consultation patterns, such as ethnicity and employment status. Using GP recorded data, it is likely that regular users are over-represented in the analysis.

For peer review only

## Introduction

Average life expectancy is shorter for men than for women in almost all countries, but the magnitude of this female advantage varies geographically and historically, highlighting the importance of understanding what causes gender differences in health. Although biological factors<sup>1-2</sup> and health behaviours<sup>3-4</sup> may provide a partial explanation, it is very widely assumed that men and women have a different propensity to consult<sup>5</sup> and that this may be an important contributor to the gender gap in mortality.<sup>6</sup>

Many large scale studies based on survey data have reported greater use of primary health care services in women.<sup>7-11</sup> In the UK, for example, women aged 16-44 years are twice as likely as men of the same age to have visited their general practitioner (GP) in the previous 12 months.<sup>12</sup> This widely reported “female excess” in consulting has led to an assumption that women are more willing to utilise health services *in all circumstances* and at all ages. However, existing evidence on gender and health service utilisation is mixed,<sup>13-17</sup> and evidence comparing consultation patterns in men and women with similar morbidity is surprisingly sparse and weak.<sup>18-19</sup> Qualitative research has indeed identified that men commonly express a reluctance to consult,<sup>20-21</sup> reinforcing a presumption that men may present with serious disease at a later (and less treatable) stage. Although very few gender comparative studies have been done in this area, there is evidence first that both men *and* women with high levels of multiple morbidity express reluctance to consult or over-use primary care resources, **in the UK at least**,<sup>22</sup> and secondly, that some groups of men express a *readiness* to consult.<sup>23-24</sup> As populations live longer with increasing health service needs, it is important to understand whether there are population subgroups who ‘under-’ or ‘over-consult’ to ensure the most effective use of primary care resources.

In the UK, general practice is usually the first point of access to formal health services, and about 90% of all NHS contacts take place in general practice, with nearly 300 million consultations a year.<sup>25</sup> There has been an upward trend in GP consultation rates between 1995 and 2008, with higher rates recorded for women than men, except in those under 15 years and over 80 years.<sup>7</sup> A similar pattern has been reported in other countries.<sup>26-28</sup> This may reflect greater medicalisation of women’s lives during the reproductive years, gender differences in underlying morbidity, or a different propensity to consult about some kinds of symptoms.<sup>15-16 29</sup> Despite the apparently consistent evidence for higher use of primary health care services in women, two important limitations to the evidence base remain. First, studies rely largely on self-reported survey data, and secondly there is a lack of attention to underlying morbidity. Indeed research evidence on gender and utilisation of health services which use routine data sources is surprisingly scant in the UK and elsewhere.

The aim of this study is to use routinely collected primary care data from the Health Improvement Network (THIN) to explore contemporary consultation patterns in men and women in the UK. We examine whether gender differences in consultation rates are constant across the life-course and across populations living in more or less affluent areas, and whether these differences remain after accounting for consultations related to reproduction and for two common underlying morbidities (depression and cardiovascular disease).



Methods

The Health Improvement Network (THIN) is one of the largest primary care data sources, consisting of electronic records of over ten million patients from more than 500 practices in the UK. Participating practices are broadly representative of UK general practice for patients’ demographic characteristics.<sup>30-31</sup> THIN contains anonymised patient data directly extracted from practices using the Vision general practice system. It is a clinical database which includes every consultation between a health professional and a patient. Data are held on an individual’s age, gender, registration details, clinical symptoms, medical diagnoses, laboratory tests, referrals and prescriptions. The level of consultation recording in THIN is comparable to that from other UK national data sources.<sup>32</sup> THIN also includes an area based deprivation index as a measure of patients’ socioeconomic status. Each postal enumeration district (about 150 households) is assigned a Townsend deprivation score. These areas are divided into national quintiles and patients in THIN are assigned a quintile score.

To be included in the study participating practices had to meet specific criteria, in relation to the completeness of data recording. The first criterion was that participating practices had acceptable computer use:<sup>33</sup> a practice on average continuously records at least one medical record, one additional health data record, and at least two prescriptions for each patient in each year. Secondly, practices’ reporting on mortality rates must be consistent with their patients’ demographic profile.<sup>34</sup> Patients within eligible practices entered the study on the latest of three possible dates: the date of registering with the practice; the date when the practice provided acceptable mortality rates; or the date when the practice attained an acceptable level of data recording. Patients no longer contributed data from the date they were transferred out of the practice, or were recorded as having died, or when the practice has its last data collection.

For this study we identified all direct contacts between clinicians and patients in primary care using Read codes, a hierarchical classification system that includes codes for signs and symptoms, diagnosis, procedures and investigations.<sup>35</sup> A total of 83,722 Read codes were included; these excluded 21,138 codes which concerned records for patient and practice administration, provision of services, hospital procedures and operations. Consultations for reproductive reasons were also identified: these included all consultations related to normal and abnormal pregnancy, childbirth and post-natal consultations (including complications following childbirth) and contraception including consulting records for pregnancy, childbirth and contraception were also identified. To further examine the effect of ~~common~~ underlying morbidities on consultation rates amongst men and women, we identified patients who were in receipt of medication to treat depression and cardiovascular disease as a clinical marker for patients with depression and CVD. These two conditions were selected as morbidities which: a) occur in both men and women, but there are gender differences in morbidity and mortality related to both conditions; b) are common in adult life; and c) are treated within general practice. ~~patients who were currently receiving antidepressant prescriptions or treatment for cardiovascular diseases.~~ Hence, ~~F~~for each condition we considered patients to be actively undergoing treatment for depression or cardiovascular if they had received two or more relevant prescriptions.<sup>36</sup> We chose to use prescription data rather than Read codes to identify patient with medical diagnoses of these conditions as we are aware that diagnoses are not consistently recorded by general practitioners; whilst some doctors record a diagnosis such as

depression each time a patient consults, others will not include the diagnosis on the patient record for a consultation if it has been previously recorded, whereas medications are recorded more consistently. Furthermore, the issue of a prescription suggests that the underlying morbidity is sufficiently serious to warrant medication. We used a cut-off of two or more prescriptions as this group of people were more likely to be taking their medication rather than those who had only ever received one prescription.

## Statistical Analysis

We calculated the annual consultation rate in 2010, using number of consultations recorded and the number of person years as the denominator. If a patient had more than one consultation within a day (whether face-to-face, over the telephone or home visits), we counted only one consultation for that day. We compared rates by gender, age groups and deprivation quintiles. We also calculated consultation rates which excluded all consultations for reproductive reasons (as noted earlier). In exploratory analyses (data not shown) we examined consultation rates for each of three years from 2008-2010; however the patterns were very similar and we therefore focus here on consultations in the most recent year available, 2010.

In order to examine the effect on gender differences in consultation rates of taking account of common underlying morbidities, as described above, at an individual patient level we developed generalised Poisson mixed models, including patient age as a nonlinear 5 knot restricted cubic spline, social deprivation (quintiles of Townsend scores) and gender. Practices were included as generalised random effects intercept terms. The log of the number of days followed in 2010 was included as an offset term. The model used a log link and Poisson / Gaussian error terms. We identified those who had at least one reproductive related consultation in 2010 and examined the effect of inclusion of that term or exclusion of those subjects from our models. Further we identified patients who had at least two prescriptions of drugs for CVD or depression.

Analyses were conducted in Stata 12 and SAS 9.2 X64, (SAS Institute, Cary, NC).

## Results

### Consultation rate by gender, age groups and deprivation quintiles

In 2010, there were 1,868,149 men and 1,916,898 women registered with 446 eligible practices. Practice size varied widely ranging from 732 to 29,779 (average practice size was 8487). The crude consultation rate was 32% less in men than in women: men consulted 3152 times per 1000 person years while women consulted 4607 times per 1000 person years (table 1).

The gender difference in consultation rates varied across the life course. As expected, we saw the largest difference in male and female consultation rates in the reproductive years, with a ratio of male to female consultation rate of 0.40 (95% CI 0.392-0.404) between the ages of 21 and 39 years (table 1). However, gender differences in primary care consultation rates were much narrower

amongst the youngest (under the age of 21 years, rate ratio= 0.77, 95% CI 0.760-0.780) and the oldest (over the age of 58 years, rate ratio=0.92, 95% CI 0.915-0.927). Indeed, consultation rates in men and women in the oldest age group were quite similar, respectively 6308 per 1000 person years in men and 6851 per 1000 person years in women.

Overall, consultation rates were higher amongst people living in the most deprived areas (5<sup>th</sup> quintile (3946 consultations per 1000 person years) than amongst those in the most affluent areas (1<sup>st</sup> quintile 3806 consultations per person years). However, there was a significant interaction between gender and deprivation status (accounting for gender as spline, and practices as random effects,  $p<0.001$ ); the gradient in consultation rates across deprivation quintiles was apparent in women, but not men.

**The effect of reproductive consultations and common underlying morbidities on gender differences in consultation rates**

Many more women (n=239,594) than men (n=829) had consulted for reproductive reasons, and around twice as many were receiving antidepressant medication (women=173,407; men=76,602) (Table 2); the numbers of men and women who were receiving cardiovascular medication were more similar (Table 2). Gender differences in consulting rates amongst these three groups of patients were much smaller than those seen in the general population (table 1). For example, amongst people in receipt of medication for depression ~~with depression~~, the consultation rate was 9102 per 1000 person year in men and 9961 per 1000 person year in women. The differences in consultation rates between men and women were also reduced when reproductive related consultations were excluded, although a considerable gap between male and female consultation rates remained between the ages of 15 and 60 years (figure 1).

Overall, after conditioning for age and deprivation quintile, there remained a substantial difference in consultation between men and women (relative risk (RR) 0.719 (95% CI 0.718 to 0.720, table 3)). This was in part ameliorated when reproductive consultations were accounted for (RR=0.81, 95%CI 0.809 to 0.811). When we further accounted for common underlying morbidities, the relative risks were much closer to unity: thus after accounting for being in receipt of medication for depression the RR was 0.916 (95% CI 0.913 to 0.918) and after accounting for being in receipt of medication for cardiovascular disease the RR was 0.950 (95%CI 0.948 to 0.952) (table 3). In addition, gender differences in consulting became even smaller when reproductive consultations were also accounted for amongst those on antidepressant medications (RR=0.951, 95% CI 0.948-0.954); additional adjustment for consultations for reproductive reasons had a little impact on gender differences in consulting for those with CVD.

**Discussion**

This study examined the extent of gender differences in the use of primary care services in the UK over the life-course and amongst people from different socio-economic backgrounds, before and after taking account of consultations for reproductive reasons and two common underlying morbidities, cardiovascular diseases and depression, as represented by being in receipt of

medication for these conditions). Overall, as expected, we found that men's consultation rates were over 30% lower than women's, confirming the gender pattern in primary care consultations reported in earlier self-report surveys.<sup>11</sup> The magnitude of gender differences in primary care consulting varied by age; there was very little difference in childhood and older age, and much higher rates of consulting in women than men during the reproductive and mid-life years. The variation in gender differences over the life-course which we observed is very similar to that reported in another study using an alternative source of routinely collected data on primary care consultations in the UK.<sup>7</sup>

These studies together provide strong evidence that, on average during the working years men have fewer contacts with their GPs than women do. It is often suggested that women's higher rates of health service contacts can be attributed to consulting for reproductive health,<sup>37</sup> but in these data consultations for reproductive reasons only partially explained the large gap in consulting between men and women in mid-life, reflecting findings from small-scale studies using self-reported data.<sup>5 38</sup> The gender difference in consulting also varied by deprivation status, reflecting a socio-economic gradient in consulting rates amongst women but not men. This finding was unexpected and warrants further exploration.

Another important finding of the study is that men and women with similar underlying morbidities (depression and cardiovascular disease) differ much less in their use of primary health care services than men and women in the population as a whole. For this analysis, we choose two common conditions, which are usually managed by general practitioners, both of which differ by gender for morbidity and mortality. Compared to an overall male to female consultation rate ratio of 0.72, the ratio was 0.92 for patients with-in receipt of medication for depression and 0.95 for patients with-in receipt of medication for cardiovascular disease. This mirrors findings from smaller studies which have relied on self-reported data on morbidity and consulting which suggest that gender differences in consulting are small after taking account of underlying morbidity.<sup>19 39</sup> Our findings suggest that some of the "excess" in female consulting may reflect greater levels of 'need' in relation to depression in women. However, given that we identified patients with depression through their medication records, it is important to note that previous studies have reported that women are more likely to receive a prescription when consulting their GPs.<sup>40</sup> Well-designed experimental studies in the UK and the USA have shown how GPs take account of gender, ethnicity and age in assessing the likely importance of symptom presentations: given the same presentations of symptoms of coronary heart disease (CHD), GPs were more likely to attribute these to CHD in men and to have a higher level of certainty about their diagnosis.<sup>41-43</sup> Although gender comparative evidence on consulting for other potentially fatal diseases is sparse,<sup>44</sup> a US study which compared the consulting histories of men and women diagnosed with colorectal cancer found that on average women delayed *longer* than men after first noticing their symptoms, and made more visits to the doctor before gaining their diagnosis. A more recent UK study of a consecutive series of lung cancer patients at three Scottish hospitals found that gender was not a predictor of time from first noticing symptoms to consulting a general practitioner.<sup>45</sup>

The strengths of this study include the use of a UK wide primary care database and the large study population which is representative of the UK population, enabling a national picture of consultation pattern. However, there are also limitations which should be noted. The aim of this study was to explore the effect of gender on the use of primary care services after considering other confounding characteristics. We were not able to examine to what extent gender differences in consultation

pattern are influenced by factors such as ethnicity and employment status because of limited completeness of data recording on these variables. Studies have shown the effect of ethnicity and gender concordance (similarity in gender of doctor and patients) on the use of services.<sup>46-47</sup> Secondly, our analyses focused on the utilisation of primary care health services, but were not able to examine fully the level of clinical needs. Thirdly, THIN includes GP recorded data, which cover consultation information about patients who have been in contact with GPs or nurses, and it is likely that regular users are over-represented in the analysis. Finally, as noted earlier, we included patients who had 2 two or more prescriptions as a proxy for underlying morbidity, but this may have failed to identify patients with milder depression/ cardiovascular disease if their condition was not severe enough to warrant medication. In addition, some prescriptions such as antidepressants can be used to treat other conditions, such as chronic pain. Whilst this is likely to occur relatively infrequently, if there are gender differences in these other conditions, this could differentially affect the specificity of controlling for use of that medication in relation to gender differences in consulting for the conditions of interest. Furthermore, if there is any difference in the compliance with medication by gender, with women being more likely to pick up a second prescription, then this could explain some of the diminution in gender differences in consulting when controlling for medication use.

Despite these limitations, this study provides much needed information on recent primary care use in men and women. It shows that gender differences in consulting are not universal; indeed the magnitude of gender differences are modest at some stages of the life course, and when account is taken of underlying morbidity.

**Conclusions**

Differences in consultations between men and women are most marked between the ages of 16 and 60 years, confirming that on average men have fewer contacts with general practitioners in early adulthood and mid-life, and these differences are that is only partially accounted for when consultations for reproductive health are considered. However, gender differences in consultations rates in patients with in receipt of medication for cardiovascular disease and depression are relatively small, suggesting men and women with common morbidities may have more similar patterns of consulting. General practitioners need to be aware in planning their delivery of health care that the gender difference in primary care health service utilisation are not constant and do not simply reflect a greater and universal propensity for women to consult more readily than men.

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**Table 1:** Crude consultation rate per 1000 person year in 2010, by gender, age groups and deprivation

	Men					Women					M/F ratio (95% CI)
	No.	/	person year	Rate	(95% CI)	No.	/	person year	Rate	(95% CI)	
<b>Overall</b>	5361100		1700883	<b>3152</b>	(3149 3155)	8001121		1736618	<b>4607</b>	(4604 4610)	<b>0.68</b> (0.68 0.69)
<b>Age quartiles</b>											
0-20	833499		415772.6	<b>2005</b>	(2000 2009)	1033997		396957.6	<b>2605</b>	(2600 2610)	<b>0.77</b> (0.76 0.78)
21-39	705785		425352.9	<b>1659</b>	(1655 1663)	1766972		423892.6	<b>4168</b>	(4162 4174)	<b>0.40</b> (0.39 0.40)
40-57	1225957		448215.3	<b>2735</b>	(2730 2740)	1914647		436238.1	<b>4389</b>	(4383 4395)	<b>0.62</b> (0.62 0.63)
58+	2595859		411542.2	<b>6308</b>	(6300 6315)	3285505		479529.5	<b>6851</b>	(6844 6859)	<b>0.92</b> (0.91 0.93)
<b>Deprivation quintiles<sup>1</sup></b>											
1	1380470		432255.4	<b>3194</b>	(3188 3199)	1940611		440269.4	<b>4408</b>	(4401 4414)	<b>0.72</b> (0.72 0.73)
2	1176355		363717.9	<b>3234</b>	(3228 3240)	1699076		373145.0	<b>4553</b>	(4546 4560)	<b>0.71</b> (0.70 0.72)
3	1073696		343302.0	<b>3127</b>	(3122 3133)	1633282		351722.6	<b>4644</b>	(4636 4651)	<b>0.67</b> (0.67 0.68)
4	940676		302437.5	<b>3110</b>	(3104 3117)	1478517		308806.1	<b>4788</b>	(4780 4795)	<b>0.65</b> (0.64 0.66)
5	655705		212989.0	<b>3078</b>	(3071 3086)	1018719		211267.7	<b>4822</b>	(4812 4831)	<b>0.64</b> (0.63 0.65)

1. 114,537 patients' socio-deprivation data were missing.

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**Table 2:** Number of patients and crude consultation rates amongst 3 groups of patients in 2010

	Men	Women	All
No. of patients who had at least one consultation for reproductive reasons (%)	829 (0.04)	239,594 (12.50)	240,423(6.35)
age (median years)	37.5	29.5	29.5
crude consultation rate (95% CI)	5999 (5832 6169)	6283 (6272 6293)	6282 (6272 6292)
No. of patients in <u>receipt of medication for with CVD</u> (%)	221,734 (11.87)	254,831 (13.29)	476,565 (12.59)
age (median years)	66.5	68.5	67.5
crude consultation rate (95%CI)	9441 (9429 9454)	10180 (10167 10192)	9836 (9827 9845)
No. of patients in <u>receipt of medication for with depression</u> (%)	76,602 (4.10)	173,407 (9.05)	250,009 (6.61)
age (median years)	51.5	50.5	50.5
crude consultation rate (95%CI)	9102 (9081 9124)	9961 (9946 9976)	9698 (9686 9711)

**Table 3:** Gender differences in consultations adjusted for age and deprivation

	Relative Risk (95 % CI)	No. Males (%)	No. Females (%)
All patients	0.719 (0.718 0.720)	1868149 (49.36)	1916898 (50.64)
Patients conditioning for consulting for reproductive reasons	0.810 (0.809 0.811)	1868149 (49.36)	1916898 (50.64)
Patients <u>in receipt of medication for</u> CVD	0.950 (0.948 0.952)	221734 (46.53)	254831 (53.47)
Patients <u>in receipt of medication for with</u> depression	0.916 (0.913 0.918)	76602 (30.64)	173407 (69.36)
Patients <u>in receipt of medication for with</u> CVD and adjusted for reproductive reasons	0.957 (0.955 0.959)	221734 (46.53)	254831 (53.47)
Patients <u>in receipt of medication for with</u> depression and adjusted for reproductive reasons	0.950 (0.947 0.953)	76602 (30.64)	173407 (69.36)

STROBE Statement—Checklist of items that should be included in reports of *cross-sectional studies*

	Item No	Recommendation
<b>Title and abstract</b>	1	(a) Indicate the study’s design with a commonly used term in the title or the abstract (b) Provide in the abstract an informative and balanced summary of what was done and what was found
<b>Introduction</b>		
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported
Objectives	3	State specific objectives, including any prespecified hypotheses
<b>Methods</b>		
Study design	4	Present key elements of study design early in the paper
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection
Participants	6	(a) Give the eligibility criteria, and the sources and methods of selection of participants
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable
Data sources/measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group
Bias	9	Describe any efforts to address potential sources of bias
Study size	10	Explain how the study size was arrived at
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why
Statistical methods	12	(a) Describe all statistical methods, including those used to control for confounding (b) Describe any methods used to examine subgroups and interactions (c) Explain how missing data were addressed (d) If applicable, describe analytical methods taking account of sampling strategy (e) Describe any sensitivity analyses
<b>Results</b>		
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed (b) Give reasons for non-participation at each stage (c) Consider use of a flow diagram
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders (b) Indicate number of participants with missing data for each variable of interest
Outcome data	15*	Report numbers of outcome events or summary measures
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included (b) Report category boundaries when continuous variables were categorized (c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses

<b>Discussion</b>		
Key results	18	Summarise key results with reference to study objectives
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence
Generalisability	21	Discuss the generalisability (external validity) of the study results
<b>Other information</b>		
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based

\*Give information separately for exposed and unexposed groups.

**Note:** An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at <http://www.plosmedicine.org/>, Annals of Internal Medicine at <http://www.annals.org/>, and Epidemiology at <http://www.epidem.com/>). Information on the STROBE Initiative is available at [www.strobe-statement.org](http://www.strobe-statement.org).